

SOLUCIONES ALS PROBLEMES PROPOSATS AL VOLUM 23 N. 2

PROBLEMA N. 76

1. De $\Pi_n = \frac{\lambda_0 \lambda_1 \cdots \lambda_{n-1}}{\mu_1 \mu_2 \cdots \mu_n} \Pi_0$, al ser $\lambda_n = \lambda, \mu_n = \mu, \forall n \Rightarrow$

$$\Pi_n = \Psi^n \Pi_0, \quad \psi = \frac{\lambda}{\mu}$$

$$1 = \sum_{n=0}^{\infty} \Pi_n = \frac{1}{1-\psi} \Pi_0 \Rightarrow \Pi_0 = 1-\psi \Rightarrow \boxed{\Pi_n = \psi^n (1-\psi) \forall n}$$

2. La función generatriz de probabilidades resulta:

$$G(s) = \sum_{n_0}^{\infty} s^n \psi^n (1-\psi) = (1-\psi) \sum_{n=0}^{\infty} (s\psi)^n = (1-\psi)(1-s\psi)^{-1}$$

El número medio de clientes en el sistema será por tanto:

$$L = G'(s) \Big|_{s=1} = -\psi(1-\psi)(1-s\psi)^{-2} \Big|_{s=1} = \frac{\psi}{1-\psi}$$

El número medio de *estaciones desocupadas* será: $\emptyset = 1 \times \Pi_0 = 1 - \psi$

$$L = L_q + s - \emptyset \Rightarrow$$

El número medio de clientes esperando: $L_q = \frac{\psi}{1-\psi} - 1 + (1-\psi) = \frac{\psi^2}{1-\psi}$

3. $P(X \leq n) = \sum_{i=0}^n \psi^i (1-\psi) = (1-\psi) \frac{1-\psi^{n+1}}{1-\psi} = 1 - \psi^{n+1}$

La probabilidad de espera: $P(X > s = 1) = 1 - \Pi_0 = 1 - (1-\psi) = \psi$.

El tiempo medio de espera: $T_q = \frac{L_q}{\lambda} = \frac{\psi}{\mu - \lambda}$

¹La existencia de régimen permanente supone que $\psi < 1$.

4. Como $\psi < 1 \Rightarrow \Pi_0$ es la mayor probabilidad $\equiv n = 0$ es el estado más probable.

5. $\Pi'_n(\psi) = -\psi^n + (1 - \psi)n\psi^{n-1};$

$\Pi'(\psi) = 0 \Rightarrow \psi = (1 - \psi)n \Rightarrow 1/n = 1/\psi - 1 \Rightarrow$

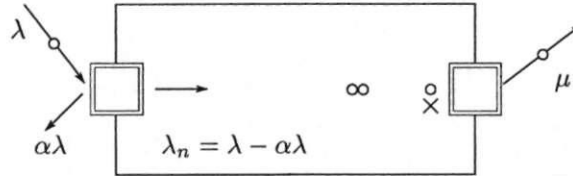
$$1/n + 1 = \frac{1+n}{n} = 1/\psi \Rightarrow \boxed{\psi = \frac{n}{n+1}} \quad \begin{array}{l} n = 1, \psi = 1/2 = 3/6 \\ n = 2, \psi = 2/3 = 4/6 > 3/6 \end{array}$$

$$\max_{\psi} \Pi_n(\psi) = \left(\frac{n}{n+1}\right)^n \left(1 - \frac{n}{n+1}\right)^{1/n+1} = \frac{n^n}{(n+1)^{n+1}} = \begin{cases} 1/4 & n = 1 \\ 4/9 & n = 2 \end{cases}$$

R. Alonso

PROBLEMA N. 77

a) Las tasas de *entrada* son ahora $\lambda_0 = \lambda$; $\lambda_n = \lambda(1 - \alpha)$, $n \geq 1$.



$$1. \Pi_1 = \frac{\lambda}{\mu} \Pi_0 = \psi \Pi_0, \quad \Pi_2 = \frac{\lambda \lambda (1 - \alpha)}{\mu \mu} \Pi_0 = \psi^2 (1 - \alpha) \Pi_0, \dots, \Pi_n = \psi^n (1 - \alpha) \Pi_{n-1} = \psi^n (1 - \alpha)^{n-1} \Pi_0 \quad \text{Nota: } \psi(1 - \alpha) < 1.$$

$$1 = \Pi_0 + \sum_{n=1}^{\infty} \Pi_n = \Pi_0 + \frac{\psi \Pi_0}{1 - \psi(1 - \alpha)} = \frac{1 + \alpha \psi}{1 - \psi(1 - \alpha)} \Pi_0 \Rightarrow$$

$$\Pi_0 = \frac{1 - (1 - \alpha)\psi}{1 + \alpha \psi} \Rightarrow$$

$$\boxed{\Pi_n = \psi^n (1 - \alpha)^{n-1} \Pi_0 = \psi^n (1 - \alpha)^{n-1} \frac{1 - (1 - \alpha)\psi}{1 + \alpha \psi} \quad \forall n > 0} \quad \alpha = 0 \quad \psi^n (1 - \psi)$$

$$2. G(s) = \Pi_0 + \sum_{n=1}^{\infty} s^n (\psi^n (1 - \alpha)^{n-1}) \Pi_0 = \Pi_0 \left(1 + s\psi \sum_{n=1}^{\infty} (s\psi(1 - \alpha))^{n-1} \right) = \Pi_0 \left(1 + s\psi \frac{1}{1 - s\psi(1 - \alpha)} \right)$$

$$G'(s) = \Pi_0 \left(\psi \frac{1 - s\psi(1 - \alpha) + s(\psi(1 - \alpha))}{(1 - s\psi(1 - \alpha))^2} \right) \Rightarrow$$

$$G'(1) = \frac{1(1 - \alpha)\psi}{1 + \alpha\psi} \left(\psi \frac{1}{(1 - \psi(1 - \alpha))^2} \right) \Rightarrow$$

$$L = \frac{\psi}{(1 + \alpha\psi)(1 - \psi(1 - \alpha))} = \begin{cases} \frac{\psi}{1 - \psi} & \alpha = 0 \\ \frac{\psi}{1 + \psi} & \alpha = 1 \end{cases}$$

$$\emptyset = \Pi_0$$

$$L_q = L - s + \emptyset = L - (1 - \Pi_0) = \frac{\psi}{(1 + \alpha\psi)(1 - \psi(1 - \alpha))} - \frac{\psi}{1 + \alpha\psi} = \frac{\psi}{1 + \alpha\psi} \left(\frac{\psi(1 - \alpha)}{(1 - \psi(1 - \alpha))} \right) \stackrel{\alpha=0}{=} \frac{\psi^2}{1 - \psi}$$

$$3. P(X > s = 1) = 1 - \Pi_0 = 1 - \frac{1 - (1 - \alpha)\psi}{1 + \alpha\psi} = \frac{\psi}{1 + \alpha\psi} \stackrel{\alpha=0}{=} \psi$$

$$\begin{aligned} \bar{\lambda} &= \lambda \Pi_0 + \sum_{n=1}^{\infty} \lambda(1 - \alpha)\Pi_n = \lambda \left(\Pi_0 + (1 - \alpha) \sum_{n=1}^{\infty} \Pi_n \right) = \\ &= \lambda (\Pi_0 + (1 - \alpha)(1 - \Pi_0)) = \lambda(1 - \alpha(1 - \Pi_0)) = \\ &= \lambda \left(1 - \alpha \frac{\psi}{1 + \alpha\psi} \right) = \frac{\lambda}{1 + \alpha\psi} \end{aligned}$$

$$T_q = \frac{L_q}{\bar{\lambda}} = \frac{1}{\lambda} \frac{\psi^2(1 - \alpha)}{(1 - \psi(1 - \alpha))} \stackrel{\alpha=0}{=} \frac{\lambda}{\mu - \lambda}$$

b. Si $\alpha = 1$ el rechazo a entrar cuando la unidad está ocupada es total:

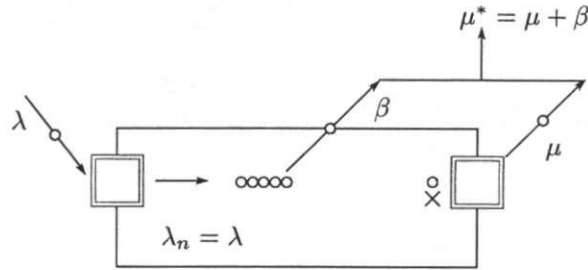
$$\lambda_n = \begin{cases} \lambda & n = 0 \\ 0 & n \geq 1 \end{cases} \Rightarrow \Pi_1 = \frac{\lambda}{\mu} \Pi_0, \Pi_n = \frac{0}{\mu} \Pi_0 = 0 \forall n > 1 \Rightarrow$$

$$\Psi \Pi_0 + \Pi_0 = 1 \Rightarrow \boxed{\Pi_0 = \frac{1}{1 + \psi}, \Pi_1 = \frac{\psi}{1 + \psi}} \quad \boxed{\emptyset = 1 \Pi_0 = \frac{1}{1 + \psi}}$$

$$\boxed{L = 0 \Pi_0 + 1 \Pi_1 = \Pi_1 = \frac{\psi}{1 + \psi}} = \begin{matrix} L_q & + & s & - & \emptyset \\ 0 & & 1 & & \frac{1}{1 + \psi} \end{matrix}$$

R. Alonso

PROBLEMA N. 78



$$1. \Pi_n = \frac{\lambda^n}{\mu(\mu + \beta)^{n-1}} \Pi_0 \quad \forall n > 1$$

$$\left(1 + \sum_{n=1}^{\infty} \frac{\lambda^n}{\mu(\mu + \beta)^{n-1}} \right) \Pi_0 = 1 \Rightarrow \Pi_0 = \frac{\mu(\mu + \beta - \lambda)}{\mu^2 + (\lambda + \mu)\beta}$$

$$\Pi_0^{-1} = 1 + \frac{\lambda}{\mu} \frac{1}{1 - \frac{\lambda}{\mu + \beta}} = 1 + \frac{\lambda}{\mu} \frac{\mu + \beta}{\mu + \beta - \lambda} = \frac{\mu^2 + (\lambda + \mu)\beta}{\mu(\mu + \beta - \lambda)}$$

$$G(s) = \Pi_0 + \sum_{n=1}^{\infty} s^n \frac{\lambda^n}{\mu(\mu + \beta)^{n-1}} \Pi_0 = \Pi_0 \left\{ 1 + s\psi \sum_{n=1}^{\infty} \left(\frac{s\lambda}{\mu + \beta} \right)^{n-1} \right\} =$$

$$= \Pi_0 \left\{ 1 + s\psi \frac{1}{1 - \frac{s\lambda}{\mu + \beta}} \right\} = \Pi_0 \left\{ 1 + \psi(\mu + \beta) \frac{s}{\mu + \beta - s\lambda} \right\} \Rightarrow$$

$$G'(s) = \Pi_0 \left(\psi(\mu + \beta) \frac{(\mu + \beta - s\lambda) - s(-\lambda)}{(\mu + \beta - s\lambda)^2} \right)$$

$$2. L = G'(1) = \frac{\mu(\mu + \beta - \lambda)}{\mu^2 + (\lambda + \mu)\beta} \left(\frac{\lambda}{\mu} / (\mu + \beta) \frac{\mu + \beta}{(\mu + \beta - \lambda)^2} \right) =$$

$$= \frac{\lambda}{\mu^2 + (\lambda + \mu)\beta} \frac{(\mu + \beta)^2}{\mu + \beta - \lambda} \underset{\beta=0}{=} \frac{\psi}{1 - \psi}$$

R. Alonso

PROBLEMA N. 79

$$\text{a) } \Pi_n = \frac{\lambda(\lambda(1-\alpha))^{n-1}}{\mu(\mu+\alpha)^{n-1}} \Pi_0 \quad \left(1 + \sum_{n=1}^{\infty} \frac{\lambda(\lambda(1-\alpha))^{n-1}}{\mu(\mu+\alpha)^{n-1}}\right) \Pi_0 = 1$$

$$\begin{aligned} \Pi_0^{-1} &= 1 + \frac{\lambda}{\mu} \frac{1}{1 - \frac{\lambda(1-\alpha)}{\mu+\alpha}} = 1 + \frac{\lambda}{\mu} \frac{\mu+\alpha}{\mu+\alpha - \lambda(1-\alpha)} = \\ &= \frac{\mu^2 + (\lambda+\mu)\alpha + \alpha\lambda\mu}{\mu(\mu+\alpha - (1-\alpha)\lambda)} \end{aligned}$$

$$\text{b) } G(s) = \Pi_0 + \sum_{n=1}^{\infty} s^n \left(\psi \left(\frac{\lambda(1-\alpha)}{\mu+\alpha} \right)^{n-1} \right) \Pi_0 =$$

$$= \Pi_0 \left(1 + s\psi \sum_{n=1}^{\infty} \left(s \frac{\lambda(1-\alpha)}{\mu+\alpha} \right)^{n-1} \right) =$$

$$= \Pi_0 \left(1 + s\psi \frac{1}{1 - s \frac{\lambda(1-\alpha)}{\mu+\alpha}} \right) = \Pi_0 \left(1 + s\psi \frac{\mu+\alpha}{\mu+\alpha - s\lambda(1-\alpha)} \right)$$

$$G'(s) = \Pi_0 \left(\psi(\mu+\alpha) \frac{\mu+\alpha - s\lambda(1-\alpha) + s\lambda(1-\alpha)}{(\mu+\alpha - s\lambda(1-\alpha))^2} \right)$$

$$L = G'(1) = \frac{\mu(\mu+\alpha - (1-\alpha)\lambda)}{\mu^2 + (\lambda+\mu)\alpha + \alpha\lambda\mu} \left(\frac{\lambda}{\mu} (\mu+\alpha) \frac{\mu+\alpha}{(\mu+\alpha - \lambda(1-\alpha))^2} \right) =$$

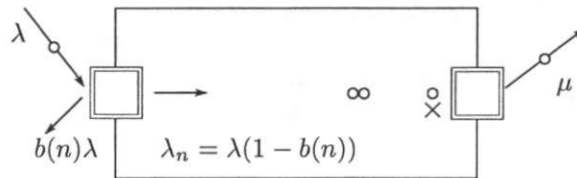
$$= \frac{\lambda}{\mu^2 + (\lambda+\mu)\alpha + \alpha\lambda\mu} \left(\frac{(\mu+\alpha)^2}{(\mu+\alpha - \lambda(1-\alpha))} \right) =$$

$$= \begin{cases} \frac{\lambda}{\mu^2 + (\lambda+\mu)\alpha} \left(\frac{(\mu-\beta)^2}{\mu+\beta-\lambda} \right), & \alpha = 0 \\ \frac{\psi}{(1+\alpha\psi)(1-\psi(1-\alpha))}, & \beta = 0 \end{cases}$$

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PROBLEMA N. 80

Las tasas de *entrada* son ahora: $\lambda_0 = \lambda$; $\lambda_n = \lambda(1 - \alpha).n \geq 1$.



$$\begin{aligned} \text{a) } \lambda_n &= \left(1 - \frac{n}{1+n}\right) \lambda = \frac{1}{1+n} \lambda \Rightarrow \Pi_n = \frac{\frac{1}{1} \lambda \frac{1}{2} \lambda \cdots \frac{1}{n} \lambda}{\mu \mu \cdots \mu} \Pi_0 = \\ &= \frac{\psi^n}{n!} \Pi_0 \Rightarrow \Pi_0 = e^{-\psi} \Rightarrow \Pi_n = \frac{\psi^n}{n!} e^{-\psi} \equiv X \in \mathcal{P}(\psi) \end{aligned}$$

$$\text{b) } L = E(X) = \psi \quad L_q = L - s + \emptyset = L - (1 - \Pi_0) = \psi - 1 + e^{-\psi}$$

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PROBLEMA N. 81

a) $\Pi_1 = 1 \Pi_0, \Pi_2 = 1(\Pi_1 = \Pi_0), \dots, \Pi_n = \Pi_0, \forall n \leq K;$

$$1 = \sum_{n=0}^K \Pi_0 \Rightarrow \boxed{\Pi_n = \frac{1}{1+K}, \quad n \leq K}$$

b) $L = \sum_{n=0}^K n \Pi_0 = \frac{1}{1+K} \frac{K(K+1)}{2} = \frac{K}{2}$

c) $\emptyset = 1 \times \Pi_0 = \frac{1}{1+K}$

d) $L_q = L - s + \emptyset = \frac{K}{2} - 1 + \frac{1}{1+K} = \frac{K(1+K) - 2(1+K) + 2}{2(1+K)} \stackrel{2}{=} \frac{K(K-1)}{2(1+K)}$

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$${}^2L_q = \sum_{n=2}^K (n-1)\Pi_0 = \frac{1}{1+K} \sum_{n=2}^K (n-1) = \frac{1}{1+K} \sum_{n=1}^{K-1} n = \frac{1}{1+K} \frac{(K-1)K}{2}$$

PROBLEMES PROPOSATS

PROBLEMA N. 82

Let $\mathbf{x}_1, \dots, \mathbf{x}_n$ independently distributed with $\mathbf{x}_i \sim N_p(\mu_i, \Omega)$, where Ω is nonsingular. It is assumed that $n > p$ and that the matrix $M' = (\mu_1, \dots, \mu_n)$ is of rank 1. Define

$$S = \sum_{i=1}^n \mathbf{x}_i \mathbf{x}_i'$$

Find $E(S^{-1})$ as a second order approximation to the exact solution of Steerneman (1997, 1999), given by

$$E(S^{-1}) = \frac{1}{n - (p + 1)} \Omega^{-1}.$$

(The author invite readers to propose a solution).

References

Steerneman, A.G.M. (1997). «Problem 331», *Statistica. Neerlandika*, 51, 381.

Steerneman, A.G.M. (1999). «Solution 331», *Statistica. Neerlandika*, 53, 252-254.

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PROBLEMA N. 83

El estimador de regresión lineal multivariante de la media de una población finita de tamaño N , $\bar{Y} = (1/N) \sum_{i=1}^N y_i$, usando el principio de mínimos cuadrados en base a n observaciones de la variable de interés y , siendo x_j ($j = 1, 2, \dots, k$) la j -ésima variable auxiliar, es:

$$\hat{Y} = \bar{y} + \sum_{j=1}^k b_j (\bar{X}_j - \bar{x}_j),$$

donde \bar{y} es la media muestral de la variable de interés, b_j el coeficiente de regresión j -ésimo, y \bar{x}_j y \bar{X}_j son la media muestral y poblacional respectivamente para la j -ésima variable auxiliar. Demostrar que si los coeficientes de regresión están acotados, el estimador \hat{Y} es consistente.

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