

PROBLEMA N. 82

Let x_1, \dots, x_n independently distributed with $x_i \sim N_p(\mu_i, \Omega)$, where Ω is nonsingular. It is assumed that $n > p$ and that the matrix $M' = (\mu_1, \dots, \mu_n)$ is of rank 1. Define

$$S = \sum_{i=1}^n x_i x_i'$$

Find $E(S^{-1})$ as a second order approximation to the exact solution of Steerneman (1997, 1999), given by

$$E(S^{-1}) = \frac{1}{n - (p + 1)} \Omega^{-1} - \frac{1}{n - (p + 1)} \left\{ \sum_{j=0}^{\infty} \frac{e^{-\frac{1}{2}\lambda} (\frac{1}{2}\lambda)^j}{j!} \cdot \frac{1}{n + 2j} \right\} \Omega^{-1} M' M \Omega^{-1},$$

where $\lambda := \text{tr} \Omega^{-1} M' M$.

(The author invites readers to propose an alternative solution).

References

Steerneman, A.G.M. (1997). «Problem 331», *Statistica. Neerlandica*, 51, 381.

Steerneman, A.G.M. (1999). «Solution 331», *Statistica. Neerlandica*, 53, 252-254.

Heinz Neudecker
Cesaro, Schagen
The Netherlands
heinz@fee.uva.nl

SOLUCIONS ALS PROBLEMES PROPOSATS AL VOLUM 24 N. 1

PROBLEMA N. 82

It is well known that $S \sim W_p(n, \Omega, \Omega^{-1}M'M)$, i.e. S follows a non-central Wishart distribution with scale matrix Ω and non-centrality matrix $\Omega^{-1}M'M$. As M' is of rank 1 we can write $M' = \mu\ell'$, hence $M'M = \lambda\mu\mu'$ where $\lambda := \ell'\ell$.

Define $\tilde{\Omega} := n\Omega + \lambda\mu\mu'$.

It follows that

$$(1) \quad E(S) = \tilde{\Omega}$$

$$(2) \quad D(\text{vec}S) = \frac{1}{n} (I_{p^2} + K_{pp}) \left(\tilde{\Omega} \otimes \tilde{\Omega} - \lambda^2 \mu\mu' \otimes \mu\mu' \right).$$

See, e.g. Magnus and Neudecker (1979) for these two results and related concepts like the commutation matrix K_{pp} , the vec-operator and the Kronecker product.

We shall now prove the following

$$\textit{Theorem: } E(S^{-1}) \approx \frac{n + (p+1)}{n} \tilde{\Omega}^{-1} - \frac{2\lambda^2}{n} \left(\mu' \tilde{\Omega}^{-1} \mu \right) \tilde{\Omega}^{-1} \mu\mu' \tilde{\Omega}^{-1}.$$

Proof: Consider the $p \times p$ continuously differentiable symmetric matrix function Z^{-1} and expand this in the point $Z = \tilde{\Omega}$, with $dZ = S - \tilde{\Omega}$ and $\Delta Z^{-1} = S^{-1} - \tilde{\Omega}^{-1}$, viz

$$S^{-1} \approx \tilde{\Omega}^{-1} - \tilde{\Omega}^{-1} (S - \tilde{\Omega}) \tilde{\Omega}^{-1} + \tilde{\Omega}^{-1} (S - \tilde{\Omega}) \tilde{\Omega}^{-1} (S - \tilde{\Omega}) \tilde{\Omega}^{-1}.$$

We used $\Delta Z^{-1} \approx dZ^{-1} + \frac{1}{2} d^2 Z^{-1} = -Z^{-1}(dZ)Z^{-1} + Z^{-1}(dZ)Z^{-1}(dZ)Z^{-1}$.

For matrix differentials see Magnus and Neudecker (1999).

Then

$$E(S^{-1}) \approx \tilde{\Omega}^{-1} + E \left\{ \tilde{\Omega}^{-1} (S - \tilde{\Omega}) \tilde{\Omega}^{-1} (S - \tilde{\Omega}) \tilde{\Omega}^{-1} \right\}, \quad \text{as}$$

$$E(S - \tilde{\Omega}) = 0.$$

Vectorization of $E \left\{ (S - \tilde{\Omega}) \tilde{\Omega}^{-1} (S - \tilde{\Omega}) \right\}$ yields

$$\begin{aligned}
\text{vec} E \left\{ (S - \tilde{\Omega}) \tilde{\Omega}^{-1} (S - \tilde{\Omega}) \right\} &= E \left\{ (S - \tilde{\Omega}) \otimes (S - \tilde{\Omega}) \right\} \text{vec} \tilde{\Omega}^{-1} \\
&= \text{vec} \left[E \left\{ (S - \tilde{\Omega}) \otimes (S - \tilde{\Omega}) \right\} \text{vec} \tilde{\Omega}^{-1} \right] \\
&= \left(\text{vec} \tilde{\Omega}^{-1} \otimes I_{p^2} \right)' E \text{vec} \left\{ (S - \tilde{\Omega}) \otimes (S - \tilde{\Omega}) \right\} \\
&= \left(\text{vec} \tilde{\Omega}^{-1} \otimes I_{p^2} \right)' C_2^p \text{vec} E \left[\left\{ \text{vec} (S - \tilde{\Omega}) \right\} \left\{ \text{vec} (S - \tilde{\Omega}) \right\}' \right] \\
&= \left(\text{vec} \tilde{\Omega}^{-1} \otimes I_{p^2} \right)' C_2^p \text{vec} D \left\{ \text{vec} (S - \tilde{\Omega}) \right\} \\
&= \left(\text{vec} \tilde{\Omega}^{-1} \otimes I_{p^2} \right)' C_2^p \text{vec} D (\text{vec} S) \\
&= \frac{1}{n} \left(\text{vec} \tilde{\Omega}^{-1} \otimes I_{p^2} \right)' C_2^p \text{vec} \left(\tilde{\Omega} \otimes \tilde{\Omega} - \lambda^2 \mu \mu' \otimes \mu \mu' \right) + \\
&\quad + \frac{1}{n} \left(\text{vec} \tilde{\Omega}^{-1} \otimes I_{p^2} \right)' C_2^p C_3^p \text{vec} \left(\tilde{\Omega} \otimes \tilde{\Omega} - \lambda^2 \mu \mu' \otimes \mu \mu' \right),
\end{aligned}$$

where $C_2^p := I_p \otimes K_{pp} \otimes I_p$ and $C_3^p := I_{p^2} \otimes K_{pp}$.

See Neudecker and Wansbeek (1987) for these definitions and some applications.

Further

$$\begin{aligned}
&\frac{1}{n} \left(\text{vec} \tilde{\Omega}^{-1} \otimes I_{p^2} \right)' C_2^p \text{vec} \left(\tilde{\Omega} \otimes \tilde{\Omega} - \lambda^2 \mu \mu' \otimes \mu \mu' \right) \\
&= \frac{1}{n} \left(\text{vec} \tilde{\Omega}^{-1} \otimes I_{p^2} \right)' \left(\text{vec} \tilde{\Omega} \otimes \text{vec} \tilde{\Omega} - \lambda^2 \text{vec} \mu \mu' \otimes \text{vec} \mu \mu' \right) \\
&= \frac{1}{n} \left(\text{tr} \tilde{\Omega}^{-1} \tilde{\Omega} \right) \text{vec} \tilde{\Omega} - \frac{\lambda^2}{n} \left(\mu' \tilde{\Omega}^{-1} \mu \right) \text{vec} \mu \mu' \\
&= \frac{p}{n} \text{vec} \tilde{\Omega} - \frac{\lambda^2}{n} \left(\mu' \tilde{\Omega}^{-1} \mu \right) \text{vec} \mu \mu'
\end{aligned}$$

and

$$\begin{aligned}
&\frac{1}{n} \left(\text{vec} \tilde{\Omega}^{-1} \otimes I_{p^2} \right)' C_2^p C_3^p \text{vec} \left(\tilde{\Omega} \otimes \tilde{\Omega} - \lambda^2 \mu \mu' \otimes \mu \mu' \right) \\
&= \frac{1}{n} K_{pp} \left(\text{vec} \tilde{\Omega}^{-1} \otimes I_{p^2} \right)' C_3^p C_2^p C_3^p \text{vec} \left(\tilde{\Omega} \otimes \tilde{\Omega} - \lambda^2 \mu \mu' \otimes \mu \mu' \right)
\end{aligned}$$

$$\begin{aligned}
&= \frac{1}{n} K_{pp} \left(\text{vec} \tilde{\Omega}^{-1} \otimes I_{p^2} \right)' C_2^p C_3^p C_2^p \text{vec} \left(\tilde{\Omega} \otimes \tilde{\Omega} - \lambda^2 \mu \mu' \otimes \mu \mu' \right) \\
&= \frac{1}{n} K_{pp} \left(\text{vec} \tilde{\Omega}^{-1} \otimes I_{p^2} \right)' C_2^p C_3^p \left(\text{vec} \tilde{\Omega} \otimes \text{vec} \tilde{\Omega} - \lambda^2 \text{vec} \mu \mu' \otimes \text{vec} \mu \mu' \right) \\
&= \frac{1}{n} K_{pp} \left(\text{vec} \tilde{\Omega}^{-1} \otimes I_{p^2} \right)' C_2^p \left(\text{vec} \tilde{\Omega} \otimes \text{vec} \tilde{\Omega} - \lambda^2 \text{vec} \mu \mu' \otimes \text{vec} \mu \mu' \right) \\
&= \frac{1}{n} K_{pp} \left(\text{vec} \tilde{\Omega}^{-1} \otimes I_{p^2} \right)' \text{vec} \left(\tilde{\Omega} \otimes \tilde{\Omega} - \lambda^2 \mu \mu' \otimes \mu \mu' \right) \\
&= \frac{1}{n} K_{pp} \text{vec} \left\{ \left(\tilde{\Omega} \otimes \tilde{\Omega} - \lambda^2 \mu \mu' \otimes \mu \mu' \right) \text{vec} \tilde{\Omega}^{-1} \right\} \\
&= \frac{1}{n} K_{pp} \left(\tilde{\Omega} \otimes \tilde{\Omega} - \lambda^2 \mu \mu' \otimes \mu \mu' \right) \text{vec} \tilde{\Omega}^{-1} \\
&= \frac{1}{n} \left(\tilde{\Omega} \otimes \tilde{\Omega} - \lambda^2 \mu \mu' \otimes \mu \mu' \right) \text{vec} \tilde{\Omega}^{-1} \\
&= \frac{1}{n} \text{vec} \tilde{\Omega} - \frac{\lambda^2}{n} \left(\mu' \tilde{\Omega}^{-1} \mu \right) \text{vec} \mu \mu'.
\end{aligned}$$

Use was made of the property

$$C_3^p C_2^p C_3^p = C_2^p C_3^p C_2^p.$$

See Ghazal and Neudecker (1998).

Collecting the two sub-results and de-vectorizing yields

$$E \left(S - \tilde{\Omega} \right) \tilde{\Omega}^{-1} \left(S - \tilde{\Omega} \right) = \frac{p+1}{n} \tilde{\Omega} - \frac{2\lambda^2}{n} \left(\mu' \tilde{\Omega}^{-1} \mu \right) \mu \mu'.$$

Hence

$$E \left(S^{-1} \right) \approx \frac{n+(p+1)}{n} \tilde{\Omega}^{-1} - \frac{2\lambda^2}{n} \left(\mu' \tilde{\Omega}^{-1} \mu \right) \tilde{\Omega}^{-1} \mu \mu' \tilde{\Omega}^{-1}. \quad \blacksquare$$

Note

In the **centrality** case ($\mu = 0$, hence $\tilde{\Omega} = n\Omega$)

we get

$$E \left(S^{-1} \right) \approx \frac{n+(p+1)}{n^2} \Omega^{-1}.$$

The exact result is

$$E(S^{-1}) = \frac{1}{n - (p + 1)} \Omega^{-1}.$$

See, e.g., Giguère and Styan (1978). The exact and approximate expressions are close together when $n \gg p$.

Also note that the exact solution to the problem is

$$E(S^{-1}) = \frac{1}{n - (p + 1)} \Omega^{-1} - \frac{1}{n - (p + 1)} \left\{ \sum_{j=0}^{\infty} e^{-\frac{1}{2}\lambda} \frac{\left(\frac{1}{2}\lambda\right)^j}{j!} \frac{1}{n + 2j} \right\} \Omega^{-1} M' M \Omega^{-1}$$

where λ is the non zero eigenvalue of $\Omega^{-\frac{1}{2}} M' M \Omega^{-\frac{1}{2}}$, hence $\lambda = \text{tr} \Omega^{-1} M' M$.

References

- Ghazal, A.G. and H. Neudecker (1998). «On second-order and fourth-order moments of jointly distributed random matrices», submitted to *Linear Algebra and Its Applications*.
- Giguère, M.A. and G.P.H. Styan (1978). «Multivariate normal estimation with missing data on several variates». *Transactions of the Seventh Prague Conference on Information Theory, Statistical Decision Functions, Random Processes*. Academica, Prague & Reidel, Dordrecht, vol. B, 129-139.
- Magnus, J.R. and H. Neudecker (1979). «The commutation matrix: some properties and applications». *Annals of Statistics*, 7, 381-394.
- Magnus, J.R. and H. Neudecker (1999). *Matrix Differential Calculus with Applications in Statistics and Econometrics*. Revised Edition, John Wiley & Sons, Chichester.
- Neudecker, H. and T.J. Wansbeek (1987). «Fourth-order properties of normally distributed random matrices». *Linear Algebra and Its Applications*, 97, 13-24.
- Steerneman, A.G.M. (1997). «Problem 331», *Statistica. Neerlandica*, 51, 381.
- Steerneman, A.G.M. (1999). «Solution 331», *Statistica. Neerlandica*, 53, 252-254.

Heinz Neudecker
Cesaro, Schagen
The Netherlands
heinz@fee.uva.nl

PROBLEMA N. 84

1) La función de densidad de X_β , variable X condicionada a $X \leq \beta$ es

$$\begin{aligned} f_\beta(x) &= \frac{d}{dx} P(X \leq x/X \leq \beta) \\ &= \frac{d}{dx} \frac{F(x)}{F(\beta)} \\ &= \frac{f(x)}{F(\beta)} \quad 0 \leq x \leq \beta \end{aligned}$$

2) En el problema n. 49 (*Qüestió*, 17(2), 1993), se prueba que si X es una variable aleatoria con densidad $f(x)$ y

$$\Gamma_\beta = \inf_{0 \leq x \leq \beta} f(x)$$

entonces se cumple la desigualdad

$$\beta^2 \Gamma_\beta \leq \sqrt{12} \sigma$$

siendo σ la desviación típica. Aplicando esta desigualdad a la función $f_\beta(x)$ obtenemos

$$\beta^2 \inf_{0 \leq x \leq \beta} \frac{f(x)}{F(\beta)} \leq \sqrt{12} \sigma(\beta)$$

siendo $\sigma(\beta)$ la desviación típica de X_β . Obsérvese que hay igualdad si y sólo si X es uniforme en $(0, \alpha)$, con $\beta \leq \alpha$.

C.M. Cuadras
Universitat de Barcelona

PROBLEMA N. 85

1) Supongamos cierta la hipótesis nula

$$H_0 : (X, Y) \text{ tiene la misma distribución que } (Y, X).$$

Debemos probar que $P(Z \leq a) = P(-Z \leq a)$, siendo $Z = X - Y$. Tenemos:

$$\begin{aligned} P(X - Y \leq a) &= P((X, Y) \in \{(x, y) / x - y \leq a\}) \\ &= P((Y, X) \in \{(x, y) / x - y \leq a\}) && \text{(por } H_0) \\ &= P((X, Y) \in \{(x, y) / y - x \leq a\}) && \text{(intercambiando } x, y) \\ &= P(Y - X \leq a) \end{aligned}$$

Luego $Z = X - Y$ tiene la misma distribución que $-Z = Y - X$, y la distribución de Z es simétrica respecto del origen.

2) Aceptar la hipótesis

$$H_1 : \text{la mediana de } Z \text{ es positiva}$$

implica rechazar H_0 . En efecto, si H_0 es cierta, entonces

$$P(Z \leq 0) = P(Z \geq 0) = \frac{1}{2}$$

y la mediana de Z es 0. Luego H_1 es incompatible con H_0 .

3) Sea $(X_1, Y_1), \dots, (X_n, Y_n)$ una muestra aleatoria simple de (X, Y) . Si H_0 es cierta

$$P(X - Y < 0) = P(Y - X < 0) = \frac{1}{2},$$

luego el número de veces k tal que

$$X_i - Y_i > 0$$

sigue la distribución binomial $B(n, \frac{1}{2})$. El test de los signos sería un test no paramétrico adecuado para contrastar H_0 frente H_1 , pues la distribución de k , bajo H_0 , no depende de la distribución de (X, Y) . También podríamos utilizar el test del signo-rango de Wilcoxon.

C.M. Cuadras y D. Cuadras
Universitat de Barcelona

PROBLEMA PROPOSAT

PROBLEMA N. 86

Sigui $H(x, y)$ la funció de distribució bivariant del parell de variables aleatòries (X, Y) .
Siguin $F(x)$, $G(y)$ les funcions de distribució marginals. El coeficient de correlació de Spearman ρ_s és el coeficient de correlació ordinari entre $F(X)$ i $G(Y)$. És ben conegut que

$$\begin{aligned}\rho_s &= \text{corr}(F(X), G(Y)) \\ &= 12 \int_{\mathbb{R}^2} (H(x, y) - F(x)G(y)) dF(x) dG(y)\end{aligned}$$

Per tant, si X, Y són estocàsticament independents, aleshores $\rho_s = 0$, és a dir,

$$H(x, y) = F(x)G(y), \quad \forall x, y \in \mathbb{R} \quad \Rightarrow \quad \rho_s = 0.$$

Es demana provar que, en general, el recíproc no és cert: X, Y poden ser estocàsticament dependents però $\rho_s = 0$.

C.M. Cuadras
Universitat de Barcelona